

# **Effects of multi-component parenting and parental mental** health interventions on early childhood development and parent outcomes: a systematic review and meta-analysis

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## Summary

Background Interventions supporting parents of young children often target parenting or parental mental health separately. Multi-component parenting and parental mental health interventions have the potential to improve parenting practices, mental health, and early childhood development. We aimed to examine their impact on child and parent outcomes.

Methods In this systematic review and meta-analysis, we searched MEDLINE, Embase, Web of Science Core Collection, APA PsycINFO, CINAHL Complete, the Cochrane Central Register of Controlled Trials, and the Global Health Database from inception to Jan 23, 2024. Eligible studies were randomised controlled trials of interventions explicitly targeting parenting behaviours and parental mental health antenatally or in children's first 3 years of life. Screening, extraction, and quality assessment were done independently by two authors. Primary outcomes were cognitive and social-emotional functioning in children and depressive symptoms in parents, meta-analysed as standardised mean differences (SMDs), relative to control. This study is registered with PROSPERO, CRD42022302848.

Findings We found 5843 records. After screening 2636 (45.1%) titles and abstracts, we manually identified and screened three additional articles and excluded 2177 records. After screening 462 full-length articles, 25 articles, representing a sample size of 8520 children and caregivers, were included. At baseline, mean caregiver age was 27.7 years (SD 5.9) and mean child age (excluding those enrolled during pregnancy) was 14.4 months (8.0). Interventions lasted a mean of 14 months (SD 11) and used a mean of 3.7 behaviour change techniques (2.0). Most interventions dedicated more time to parenting behaviours than to parental mental health. We found significant intervention effects on children's cognitive (SMD 0.19 [95% CI 0.04 to 0.34]; P=69%) and social-emotional (0.26 [0.17 to 0.34]; P=47%) outcomes but not on depressive symptoms in female caregivers (-0.18 [-0.36 to 0.002]; I2=86%) relative to control conditions. Risk of bias across studies was moderate, and we found heterogeneity across results.

Interpretation Multi-component parenting and mental health interventions had a positive effect on child cognitive and social-emotional outcomes, but not on depressive symptoms in parents, suggesting that other factors might contribute to positive ECD outcomes. Interventions might lack adequate focus on mental health to make a discernible impact, highlighting a need for future studies to differentiate and assess contributions of parenting and mental health components to understand independent and collective effects on family outcomes.

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# Introduction

An estimated 250 million children worldwide do not meet their developmental potential.1 Parenting behaviours and parental mental health are key determinants of early childhood development (ECD) that are strongly associated with one another.2 Although many interventions to improve ECD outcomes have primarily targeted only one of these two factors, a handful of interventions have targeted both parenting behaviours and parental mental health (particularly depression), as this approach might be more effective in improving child and parent outcomes. Assessing the effect of these interventions in the first 3 years of life-when the developing brain is most sensitive to experiences and the environment<sup>3</sup>---can help to advance scientific understanding of the combined effects of improving parenting behaviours and parental mental health on enabling children to attain their full developmental potential.4

The experience of depressive symptoms in the perinatal period is a substantial global mental health challenge among mothers (11-25%) and fathers (7-11%).5,6 In addition to impairing parents' health and functioning, the experience of depressive symptoms can interfere with parents' ability to actively engage with their child and respond appropriately to their developmental needs.7 Parents experiencing depressive symptoms are more

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### **Research in context**

#### Evidence before this study

Previous systematic reviews have investigated the efficacy of parenting or parental mental health interventions in isolation, reporting mixed effects on early childhood development (ECD) and parent outcomes. Observational studies, primarily from high-income countries, have shown robust associations between parental mental health and parenting behaviours; yet, the combined impact of interventions targeting both parenting and parental mental health on parenting and child outcomes has only recently been examined. We searched MEDLINE, Embase, Web of Science Core Collection, APA PsycINFO, CINAHL Complete, the Cochrane Central Register of Controlled Trials, and the Global Health Database for studies published in English, French, Spanish, or Arabic from inception to Jan 23, 2024, using combinations of phrases for search terms such as "parenting interventions," "mental health interventions," and "child outcomes". We found no systematic reviews or meta-analyses on this topic.

### Added value of this study

To our knowledge, this study is the first to systematically analyse and quantify the effects of multi-component

interventions that simultaneously target parenting and parental mental health, offering new insights into the combined impact of these interventions on ECD and parent outcomes. Our findings indicate that although these multicomponent interventions improved children's cognitive and social-emotional outcomes, there was no significant effect on depressive symptoms in parents. Although such interventions show promise, our findings also highlight gaps for future improvement to better serve young children and their families.

# Implications of all the available evidence

The findings emphasise the importance of an integrated approach in designing interventions for parents and children. The combination of parenting and mental health components in interventions could potentially enhance the efficacy of programmes aimed at improving ECD and parent outcomes. These insights are important for informing future research, clinical practice, and policy making, emphasising the need for comprehensive strategies to improve overall family wellbeing.

likely to engage in less responsive and stimulating parenting behaviours that subsequently have negative effects on ECD outcomes.<sup>8,9</sup> Other dimensions of parental mental health (eg, parenting stress) affect parents by increasing symptoms of depression, anxiety, and fatigue, reducing the quality of parenting behaviours and increasing the risk of adverse social–emotional outcomes in children.<sup>10</sup>

The robust associations of parental mental health and parenting behaviours with ECD outcomes suggest that interventions targeting mental health or parenting behaviours should improve child and parent outcomes. However, meta-analyses of interventions to prevent, reduce, or treat depression, anxiety, or parenting stress across the perinatal period report mixed findings on their efficacy in improving ECD outcomes.<sup>11,12</sup> Similarly, metaanalytical evidence indicates that interventions in the early years focusing solely on improving parenting behaviours without addressing parents' underlying mental health are not enough to improve parental mental health outcomes.<sup>13,14</sup> These findings highlight the need for an integrated approach that jointly targets parental mental health and parenting behaviours to improve child and parent outcomes.4

Although the prevalence of depressive symptoms in mothers throughout the perinatal period in low-income and middle-income countries (LMICs; 20–25%)<sup>15</sup> is almost double that in high-income countries (11–13%),<sup>6</sup> most studies have focused on associations with child outcomes in high-income countries.<sup>16</sup> There are even fewer studies on depressive symptoms in fathers in LMICs and their

association with child outcomes.<sup>5.17</sup> Data exist on the effect of parenting interventions across sociocultural contexts, with a meta-analysis of parenting interventions finding greater effects on ECD and parenting outcomes in LMICs than in high-income countries.<sup>14</sup> Given sociocultural differences in parenting, and the availability and quality of mental health services, it is necessary to consider global evidence assessing integrated approaches to improve child and parent outcomes.

Researchers have begun to adopt this integrated approach by developing multi-component interventions that explicitly promote positive parenting behaviours and address parental mental health to improve ECD.4,18 In addition to coaching parents on the developmental importance of responsive and stimulating parent-child interactions, these interventions address the psychological wellbeing of parents by including content on topics such as emotional literacy and stress management. Despite this increase in multi-component interventions, there are no reviews of their efficacy in improving child and parent outcomes nor studies on their implementation and whether implementation features moderate intervention impact. Understanding the implementation of these multi-component interventions and their efficacy in improving child and parent outcomes is key for advancing scientific knowledge on how best to design interventions that support parents to promote their wellbeing and their child's developmental outcomes.19

The primary aim of this meta-analysis was to examine the effect of multi-component parenting and parental mental health interventions on child and parent outcomes



Figure 1: Study selection

See Online for appendix

globally. We include parent outcomes—specifically, parental mental health and parenting behaviours—as indicators of the potential mechanisms through which the interventions affect child outcomes. The secondary aim was to identify whether intervention effects on child and parent outcomes are moderated by key study characteristics and intervention implementation features.

## **Methods**

#### Search strategy and selection criteria

This systematic review and meta-analysis follows the PRISMA statement for standard reporting (appendix pp 3–5).<sup>20</sup> The protocol was preregistered on PROSPERO, CRD42022302848, and is available online (appendix pp 6–7).<sup>65</sup>

We identified articles on interventions that jointly targeted parenting behaviours and parental mental health during pregnancy and early childhood up to age 3 years by searching MEDLINE (Ovid), Embase (Elsevier), Web of Science Core Collection (Clarivate Analytics), APA PsycINFO (EBSCO), CINAHL Complete (EBSCO), the Cochrane Central Register of Controlled Trials (Wiley), and the Global Health Database (EBSCO), without any date restrictions. Our search strategy (appendix pp 8–16) was informed by search terms from relevant systematic reviews.<sup>14,21</sup> Controlled vocabulary terms (ie, Medical Subject Headings [MeSH]) were included when available. The searches, prepared by three authors (AAA, PAB, and MNA), were conducted by

PAB across the selected databases and last run on Jan 23, 2024. We also searched bibliographies of included studies and relevant reviews for further studies.

Articles were considered eligible for inclusion if they: (1) explicitly targeted interactions, behaviours, knowledge, beliefs, attitudes, or practices of parents with their children and parental mental health or wellbeing to improve ECD outcomes; (2) used a randomised controlled trial (RCT) design; (3) targeted children and their parents during early childhood (pregnancy until age 3 years);<sup>3</sup> (4) were published in a peer-reviewed journal in English, French, Spanish, or Arabic; and (5) assessed an ECD outcome and parenting practices or parental mental health after intervention completion (or shortly thereafter).

Child outcomes were two domains of ECD: (1) cognitive (ie, language, intellectual, and executive functioning capabilities) and (2) social-emotional (ie, socioemotional, behavioural, and attachment) outcomes. We focused on four parent outcomes: (1) positive parenting practices (eg, engagement in play, cognitive stimulation, and warmth), (2) negative parenting practices (eg, hostile parent-child interactions and harsh discipline), (3) depressive symptoms, and (4) parenting stress.

Two authors (AAA and MNA) independently screened titles and abstracts from the initial search. Studies deemed by both authors to not fulfil the inclusion criteria were excluded. The remaining articles were independently read by the two authors, and those meeting the inclusion criteria by mutual agreement were included. Disagreements were resolved in team discussions.

### Data analysis

Data from each study were abstracted onto a standardised form by AAA, and MNA independently extracted the same information from a subsample for quality assurance. Study characteristics, including participant race or ethnicity and sex or gender, were extracted as reported by study authors.

Data were also extracted for study quality assessment using the Cochrane Collaboration Risk of Bias Assessment Tool for RCTs.22 For each study, AAA and MNA independently rated risk of bias across seven dimensions (random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective reporting, and other) as low (0), unclear (1), or high (2). We then assigned studies an overall rating: low (low risk of bias across the first six dimensions), unclear (unclear risk of bias in at least one of these six dimensions), or high (high risk of bias in at least one of these six dimensions).22 Following Cochrane guidelines,<sup>22</sup> we accounted for differences between cluster and individual RCTs where appropriate. To examine publication bias, we used Egger's regression test and funnel plots to test the null hypothesis of smallstudy bias.23

| Parent outcomes<br>assessed                                      | Positive parenting<br>practices, depressive<br>symptoms, and<br>other (or<br>composite) mental<br>health  | Positive parenting<br>practices, depressive<br>symptoms,<br>parenting stress,<br>and other (or<br>composite) mental<br>health                         | Positive parenting<br>practices, depressive<br>symptoms, other<br>(or composite)<br>mental health | Depressive<br>symptoms  | Depressive<br>symptoms,<br>parenting stress,<br>and other (or<br>composite) mental<br>health    | Depressive<br>symptoms and<br>parenting stress   | Depressive<br>symptoms and<br>other (or<br>composite) mental<br>health<br>health<br>ntinues on next page)  |
|--|---|---|---|---|---|--|--|
| Early childhood<br>development<br>outcomes<br>assessed           | Cognitive and<br>social-<br>emotional   | Social-<br>emotional  | Cognitive and<br>social-<br>emotional   | Social-<br>emotional  | Cognitive and<br>social-<br>emotional   | Social-<br>emotional   | Social-<br>emotional<br>(Table co  |
| Post-<br>intervention<br>outcome<br>assessment,<br>months§       | 4   | 0   | 0   | 0   | 0   | 0  | 0  |
| Behaviour change<br>techniques‡                                  | Family engagement,<br>print media,<br>performance—self<br>and performance—<br>other, problem<br>solving, social<br>support—<br>support—<br>community, and<br>other (attendance<br>gift) | Print media,<br>performance—other,<br>problem solving,<br>other (therapeutic<br>alliance),<br>collaborative goal<br>setting, and rapport-<br>building | Social support —<br>community   | Family engagement   | Performance—self<br>and rapport building  | Family engagement,<br>social support—<br>community, and<br>collaborative goal<br>setting | Print media,<br>problem solving,<br>and collaborative<br>goal setting  |
| Dosage and delivery  | Four 60-min,<br>fortnightly,<br>individual home<br>visits and<br>12 90-min,<br>fortnightly,<br>community-based<br>group sessions  | 43 60-min<br>individual home<br>visits of varied<br>frequency   | 12 60–90-min,<br>weekly, individual<br>home visits  | 7 monthly<br>individual home<br>visits and<br>14 60-min,<br>fortnightly,<br>community-based<br>group sessions | 8 individual home<br>visits (twice a<br>month) and<br>14 monthly<br>individual clinic<br>visits | Individual home<br>visits (varied<br>frequency and<br>number)                            | Four 90-min<br>individual clinic<br>visits   |
| Type of<br>intervention  | Universal   | Targeted:<br>pregnant<br>adolescents<br>(American<br>Indian)  | Targeted: low-<br>income<br>families  | Universal   | Targeted:<br>families with<br>pre-term<br>infants   | Targeted: at-<br>risk families   | Targeted: at-<br>risk families   |
| Average<br>intervention<br>duration                              | 7 months  | 38 months   | 4 months  | 7 months  | 18 months   | Varied   | 6 months   |
| ✓ Mean child<br>age at<br>baseline,<br>months†                   | 14.1  | Antenatal<br>(≤32<br>weeks)   | 23·5  | Antenatal<br>(not<br>reported)  | 0   | 5.2  | Antenatal<br>(6–8<br>months)   |
| Primary beneficiaries<br>race or ethnicity and<br>sex or gender* | Luhya   | American Indian<br>100% women   | 56% female and<br>25% male¶   | Mamprusi and Nabit<br>100% women  | :   | African-American,<br>Caucasian, Latina,<br>and other 100%<br>women                       | Non-Hispanic White,<br>Hispanic White,<br>Native American or<br>Alaskan Native, Asian<br>or Pacific Islander,<br>and African American<br>50% men |
| Primary<br>beneficiaries   | Mothers and<br>fathers  | Mothers only  | Primary<br>caregivers   | Mothers<br>(fathers and<br>grandmothers<br>invited)   | Mothers and<br>fathers  | Mothers only   | Mothers and<br>fathers   |
| Sample<br>(randomised<br>controlled trial<br>type)               | Intervention<br>n=373 and<br>control m=351<br>(cluster)   | Intervention<br>n=159 and<br>control n=163<br>(individual)  | Intervention n=19<br>and control n=19<br>(individual)   | Intervention<br>n=182 and<br>control n=131<br>(cluster)   | Intervention<br>n=29 and control<br>n=24 (individual)   | Intervention<br>n=132 and control<br>n=100<br>(individual)                               | Intervention<br>n=60 and control<br>n=60 (individual)  |
|  | Luoto et al,<br>2018-19, Kenya<br>(LMIC) <sup>5</sup>   | Barlow et al,<br>2006–11, USA<br>(HIC) <sup>»</sup>   | Barnhart et al<br>(years not<br>reported),<br>Rwanda<br>(LMIC) <sup>31</sup>                      | Baumgartner<br>et al, 2018–20,<br>Ghana (LMIC) <sup>32</sup>  | Castel et al,<br>2006–10,<br>France (HIC) <sup>33</sup>   | Demeusey et al,<br>years not<br>reported, USA<br>(HIC) <sup>34</sup>                     | Doss et al, and<br>Tomfohr-<br>Madsen et al,<br>years not<br>reported, USA<br>(HIC) <sup>SS,#</sup>  |

|  | Sample<br>(randomised<br>controlled trial<br>type)          | Primary<br>beneficiaries                            | Primary beneficiaries'<br>race or ethnicity and<br>sex or gender <sup>*</sup>  | Mean child<br>age at<br>baseline,<br>months† | Average<br>intervention<br>duration | iype or<br>intervention                              | Dosage and delivery  | Behaviour change<br>techniques‡  | Post-<br>intervention<br>outcome<br>assessment,<br>months§ | Early childhood<br>development<br>outcomes<br>assessed | Parent outcomes<br>assessed   |
|--|---|---|--|--|-------------------------------------|--|--|--|--|--|---|
| ed from                                    | n previous page)  |   |  |  |                                     |  |  |  |  |  |   |
| and<br>ars not<br>I, USA                   | Intervention<br>n=114 and<br>control n=107<br>(individual)  | Mothers only  | Black or African<br>American, Latina<br>(Mexican), Multiracial<br>or Ethnic or other,<br>and White or<br>European American<br>100% women | Antenatal<br>(<34<br>weeks)                  | 30 months                           | Targeted:<br>pregnant, low-<br>income<br>adolescents | Individual home<br>visits (varied<br>frequency and<br>number)  | Materials (book and<br>age-appropriate<br>toy), performance–<br>self, rapport-<br>building, and social<br>support–<br>community        | 0  | Social-<br>emotional                                   | Positive parenting<br>practices and<br>negative parenting<br>practices  |
| n et al,<br>4, New<br>(HIC)³ <sup>8</sup>  | Intervention<br>n=184 and<br>control n= 207<br>(individual) | Primary<br>caregivers                               | White or Mãori   | 7  | 36 months                           | Targeted:<br>socially<br>vulnerable<br>families      | 176 120-min,<br>weekly, individual<br>home visits  | Collaborative goal<br>setting, problem<br>solving, social<br>support—<br>community, and<br>rapport building                            | 0  | Social-<br>emotional                                   | Negative parenting<br>practices   |
| tal,<br>1,                                 | Intervention<br>n=616 and<br>control n=544<br>(cluster)     | Mothers<br>(fathers and<br>grandmothers<br>invited) | Vietnamese or Kinh   | Antenatal<br>(<20<br>weeks)                  | 14 months                           | Universal  | 19 60–90-min<br>community-based<br>group sessions<br>every 2-4 weeks<br>One individual<br>home visit | Family engagement,<br>audiovisual and<br>print media, and<br>performance—self  | 0  | Cognitive and<br>social-<br>emotional                  | Positive parenting<br>practices and other<br>(or composite)<br>mental health  |
| urst et<br>−17,<br>a (HIC)⁴0               | Intervention<br>n=146 and<br>control n=128<br>(cluster)     | Primary<br>caregivers                               | 89% women and<br>11% men   | 25.6   | 1.5 months                          | Universal  | Six 120-min, weekly,<br>community-based<br>group sessions  | Materials (posters),<br>audiovisual and<br>print media,<br>performance—other,<br>social support—<br>community, and<br>other (homework) | 12   | Social-<br>emotional                                   | Positive parenting<br>practices, negative<br>parenting practices,<br>and other (or<br>composite) mental<br>health                 |
| t al,<br>ot<br>d,<br>ia (HIC)⁴             | Intervention<br>n=49 and control<br>n=51 (individual)       | Mothers only  | 100% women   | 8.7  | 1 day                               | Targeted:<br>socially<br>vulnerable<br>families      | One 6-h clinic visit<br>consisting of<br>individual and<br>group sessions                            | Collaborative goal<br>setting,<br>performance—self<br>and performance—<br>other, and rapport<br>building                               | 0.5  | Social-<br>emotional                                   | Depressive<br>symptoms  |
| et al,<br>2, USA                           | Intervention<br>n=358 and<br>control n=236<br>(cluster)     | Mothers only  | Non-Hispanic White,<br>and non-Hispanic<br>Black<br>100% women   | llo  | 14.7 months                         | Targeted:<br>adolescent<br>mothers                   | 118 individual<br>home visits  | Collaborative goal setting   | 0  | Social-<br>emotional                                   | Negative parenting<br>practices and<br>parenting stress   |
| et al,<br>9,<br>1                          | Intervention<br>n=541 and<br>control n=508<br>(cluster)     | Primary<br>caregivers                               | 64% female and<br>29% male**   | 21·5   | 3 months                            | Targeted: low-<br>income<br>families                 | 1275-min, weekly,<br>individual home<br>visits   | family engagement,<br>performance—self,<br>problem solving,<br>social support—<br>community  | 0  | Cognitive and<br>social-<br>emotional                  | Positive parenting practices  |
| t al,<br>ot<br>d,<br>a (HIC) <sup>44</sup> | Intervention<br>n=85 and<br>control n=69<br>(individual)    | Mothers only  | 100% women   | Antenatal<br>(not<br>reported)               | 27 months                           | Targeted: at-<br>risk families                       | 52 60–90-min,<br>weekly, individual<br>home visits   | Family engagement<br>and social support—<br>community  | 0  | Social-<br>emotional<br>(Table cor                     | Positive parenting<br>practices, depressive<br>symptoms, and<br>other (or<br>composite) mental<br>health<br>trinues on next page) |

| l Parent outcomes<br>assessed                                    |                  | Depressive<br>symptoms  | Positive parenting<br>practices and<br>depressive<br>symptoms   | Depressive<br>symptoms   | Depressive<br>symptoms,<br>parenting stress,<br>and other (or<br>composite) mental<br>health | Positive parenting<br>practices and<br>depressive<br>symptoms   | Depressive<br>symptoms and<br>other (or<br>composite) mental<br>health<br>minues on next page)                |
|--|------------------|---|---|--|--|---|---|
| Early childhood<br>development<br>outcomes<br>assessed           |                  | Cognitive and<br>social-<br>emotional   | Cognitive and<br>social-<br>emotional   | Cognitive and<br>social-<br>emotional  | Social-<br>emotional   | Cognitive   | No<br>(Table co   |
| Post-<br>intervention<br>outcome<br>assessment,<br>months§       |                  | 0   | o   | 0  | 0  | m   | 0   |
| Behaviour change<br>techniques‡                                  |                  | Family engagement,<br>performance—other,<br>problem solving,<br>rapport building,<br>and social support—<br>community | Family engagement,<br>materials (low-cost<br>toys and picture<br>books), print media,<br>performance-self<br>and performance-<br>other, problem<br>solving, and social<br>support-<br>community | Materials (play<br>materials) and<br>performance—self  | Audiovisual media  | Family engagement,<br>materials (home-<br>based play<br>materials), print<br>media,<br>performance-<br>other, problem<br>solving, rapport<br>building, social<br>support-<br>community, and<br>other (homework<br>and motivational<br>interviewing) | Collaborative goal<br>setting, family<br>engagement,<br>rapport building,<br>and social support—<br>community |
| Dosage and delivery  |                  | Ten individual home<br>visits and<br>22 community-<br>based group<br>sessions, both of<br>varied frequency            | Nine 22-5-min,<br>monthly, individual<br>home visits and<br>nine 52-5-min,<br>monthly,<br>community-based<br>group sessions   | 23 fortnightly<br>individual home<br>visits and<br>20 fortnightly<br>community-based<br>group sessions | Nine individual<br>home visits (varied<br>frequency)   | 1.5 45-min<br>individual home<br>visits and 13 75-min<br>fortnightly<br>community-based<br>group sessions   | Individual home<br>visits of varied<br>frequency  |
| Type of<br>intervention  |                  | Universal   | Universal   | Universal  | Targeted: at-<br>risk families   | Universal   | Universal   |
| Average<br>intervention<br>duration                              |                  | 30 months   | 9 months  | 12 months  | 11 months  | 7 months  | 26 months   |
| s' Mean child<br>age at<br>baseline,<br>months†                  |                  | Antenatal<br>(third<br>trimester)   | 7.3   | 8.5  | Antenatal<br>(not<br>reported)   | 22.3  | Antenatal<br>(second<br>trimester)  |
| Primary beneficiarie:<br>race or ethnicity and<br>sex or gender* |                  | 10% women   | 100% women  | Tonga  | 52% female and<br>48% male   | Lango<br>63% women and<br>37% men   | Hispanic or Latino,<br>African American,<br>White or non-<br>Hispanic, and other<br>100% women                |
| Primary<br>beneficiaries   |                  | Mothers only  | Primary<br>caregivers   | Primary<br>caregivers  | Mothers and<br>fathers   | Mothers and<br>fathers  | Mothers<br>(fathers and<br>grandmothers<br>invited)   |
| Sample<br>(randomised<br>controlled trial<br>type)               | n previous page) | Intervention<br>n=206 and<br>control n=216<br>(cluster)   | Intervention<br>n=149 and<br>control n=281<br>(cluster)   | Intervention<br>n=220 and<br>control n=215<br>(cluster)  | Intervention<br>n=34 and control<br>n=33 (individual)  | Intervention<br>n=160 and<br>control n=131<br>(duster)  | Intervention<br>n=60 and control<br>n=64 (cluster)  |
|  | (Continued from  | Maselko et al,<br>2014-19,<br>Pakistan<br>(LMIC)45  | Pitchik et al,<br>2017–18,<br>Bangladesh<br>(LMIC)*   | Rockers et al,<br>2015-16,<br>Zambia<br>(LMIC)47   | Røhder et al,<br>2017-19,<br>Denmark<br>(HIC) <sup>48</sup>                                  | Singla et al,<br>2012-13,<br>Uganda<br>(LMIC)*  | Slade et al,<br>years not<br>reported, USA<br>(HIC) <sup>30</sup>   |

We summarised study characteristics and implementation features narratively. We estimated effect sizes as the standardised mean difference (SMD) between intervention and comparison groups (eg, no intervention or standard of care) with respect to change in mean values from baseline to endline after standardisation by their pooled SD. The mean, SD, and sample size of control and intervention groups were extracted from each study to estimate the SMD. When studies did not provide these data, we contacted authors directly. Out of eight authors contacted, five provided data. In multi-arm studies, the comparison group was the intervention group without any parenting and mental health components. If intervention groups differed only in delivery modalities, one group was randomly selected for analysis. We ran sensitivity analyses to assess whether studies using this approach affected estimates. We used outcome measurements from the timepoint closest to intervention completion for analyses.

We used meta-analysis with robust variance estimation to estimate pooled effects across studies, accounting for multiple dependent effect-size estimates per trial per outcome to increase power and precision. As per our pre-registered protocol, a minimum of ten studies reporting data for a given outcome was required to run a meta-analysis for that outcome. Pooled estimates were based on random-effects models, and p values of less than 0.05 denoted statistical significance. For children's social-emotional outcomes, we meta-analysed absolute values of effect sizes because the direction of effects varied across measures, and we were interested in examining differences in magnitude rather than direction of effects. We assessed heterogeneity of pooled effect sizes using the *I*<sup>2</sup> statistic. Effect size magnitudes were interpreted in the context of public health, paediatric, and early education interventions, and the practical significance with respect to outcomes.<sup>24</sup>

We examined potential sources of heterogeneity in the pooled effect of each outcome by doing subgroup analyses using random-effects meta-regression models. We included the following pre-specified study characteristics and implementation features (the choice of which was informed by an implementation review of parenting interventions<sup>19</sup>): country income level (high-income country vs LMIC per World Bank classifications), child mean age at baseline (antenatal or <12 months vs  $\geq$ 12 months), intervention duration (<12 months vs ≥12 months), delivery modality (individual vs group and a combination of individual and group),<sup>25</sup> delivery setting (single setting [ie, home, clinic, or community] vs a combination of home, clinic, or community settings),<sup>26</sup> and study risk of bias (a binary variable based on the median split of the total risk of bias score). Additional analyses were done by nature of intervention (targeted vs universal), number of behaviour-change techniques used (<3  $\nu s \ge 3$ ), and caregiver age (based on intervention beneficiaries; continuous). The categorisation of behaviour change techniques was inspired by Aboud and Yousafzai<sup>27</sup> and Pedersen.<sup>28</sup> Two-sided p<0.05 indicated a significant difference between subgroups based on the regression of meta-analytic estimates onto each subgroup in separate models. All analyses were done using the robumeta 2.0 package<sup>29</sup> in R (version 4.0.3).

# Involvement of individuals with lived experience

We involved two individuals with lived experiences of caring for young children and living with mental health problems in all aspects of the study from the conceptualisation of research questions to manuscript preparation.

# Role of the funding source

The funder of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report.

# Results

Our search returned 5843 records. After removal of 3207 (54.9%) duplicates, we identified 2636 (45.1%) unique records; three additional records were identified through manual scanning of article references. After screening the titles and abstracts of these 2639 records, 462 (17.5%) articles underwent full-text screening and, of these, 25 (5.4%) articles published between 2005 and 2023 (24 unique RCTs) were included in the systematic review (8520 children and caregivers) and 23 in the meta-analysis (7463 children and caregivers; figure 1). Analytical sample sizes ranged from 19 to 616 individuals (mean 160 [SD 117]; table).

Mean caregiver age was 27.7 years (SD 5.9, range 18.0-39.0 years): 25.2 years (4.6) for mothers and 29.2 years (6.0) for fathers. In five (21%) studies that jointly reported fathers' and mothers' mean age or included the child's primary caregiver, the mean age was 34.3 years (4.6). Female caregivers (mothers in most cases) were the primary beneficiaries in 12 (50%) interventions, with the other 12 targeting couples (six [25%]) or the child's primary caregiver, which included a mix of mothers, fathers, and other caregivers (six [25%]), with female caregivers making up the majority of intervention beneficiaries where reported. Of the 12 interventions targeting female caregivers, three (25%) invited other caregivers (fathers or partners, grandmothers, and other family members) to some sessions. Most interventions were delivered by community volunteers (eight [33%]) and community health workers (six [25%]). Interventions' parenting content primarily focused on improving parent-child interactions and ECD outcomes through education on health, nutrition, hygiene, responsive caregiving, and early language learning, as well as implementing positive conflict resolution and discipline strategies. Authors who referenced sources for parenting interventions cited Healthy Families America and Care for Child Development. For mental health

|  | Weight<br>(%) |  | Standardised mean<br>difference (95% CI) |
|--|---------------|--|--|
| Singla et al (2015) <sup>49</sup>          | 1.19          |  | 0·37 (0·13 to 0·60)                      |
| Singla et al (2015) <sup>49</sup>          | 1.19          |  | 0·29 (0·05 to 0·52)                      |
| Castel et al (2016)33                      | 1.90          | <del></del>                                  | 0.67 (0.17 to 1.18)                      |
| Rockers et al (2016 to 2018) <sup>33</sup> | 5.28          |  | -0.11 (-0.30 to 0.08)                    |
| Rockers et al (2016 to 2018)47             | 5.28          | _  | -0.08 (-0.09 to 0.25)                    |
| Barnhart et al (2020) <sup>31</sup>        | 1.38          |  | 0·13 (-0·50 to - <mark>X</mark> ·75)     |
| Barnhart et al (2020) <sup>31</sup>        | 1.37          |  | 0.50 (-0.13 to 1.12)                     |
| Barnhart et al (2020) <sup>31</sup>        | 1.37          |  | -0.04 (-0.66 to 0.58)                    |
| Maselko et al (2020)45                     | 4.99          | -  | 0·16 (-0·03 to 0·35)                     |
| Luoto et al (2021) <sup>25</sup>           | 5.63          | +  | 0·25 (0·10 to 0·39)                      |
| Luoto et al (2021) <sup>25</sup>           | 5.63          |  | 0.21 (0.06 to 0.36)                      |
| Luoto et al (2021) <sup>25</sup>           | 5.63 -        |  | -0.13 (-0.27 to 0.02)                    |
| Jensen et al (2021) <sup>43</sup>          | 2.30 -        | <u>.                                    </u> | 0·10 (−0·34 to 0·54)                     |
| Jensen et al (2021)43                      | 2.66          |  | 0.08 (-0.31 to 0.47)                     |
| Pitchik et al (2021) <sup>46</sup>         | 4.88          |  | 0·16 (-0·03 to 0·36)                     |
| Pitchik et al (2021) <sup>46</sup>         | 4.88          |  | 0·16 (-0·04 to 0·36)                     |
| Pitchik et al (2021) <sup>46</sup>         | 4.88          | <b>-</b>                                     | 0·18 (-0·02 to 0·38)                     |
| Pitchik et al (2021) <sup>46</sup>         | 4.88          | -  | 0·14 (-0·06 to 0·34)                     |
| Fisher et al (2023) <sup>39</sup>          | 6.04          | <b>.</b>                                     | 0·16 (0·04 to 0·27)                      |
| Fisher et al (2023) <sup>39</sup>          | 6.04          | <b>-</b>                                     | 0·12 (0·01 to 0·24)                      |
| Jeong et al (2023)52                       | 4.59          | -  | 0.08 (-0.14 to 0.30)                     |
| Jeong et al (2023) <sup>52</sup>           | 4.59          | <b>-</b>                                     | 0·12 (-0·10 to 0·34)                     |
| Jeong et al (2023) <sup>52</sup>           | 4.60 -        | F  | 0.02 (-0.20 to 0.24)                     |
| Tofail et al (2023)51                      | 4.42          |  | 0.60 (0.37 to 9.84)                      |
| Tofail et al (2023) <sup>51</sup>          | 4.41          |  | 0·67 (0·44 to 0·91)                      |
| Overall (I²=69%; Cochran's Q=              | 69%)          | $\diamond$                                   | 0·19 (0·04 to 0·34)                      |
|  | -1.5 -0.5 0   | 0.5 1.5 2.5                                  | 5  |

Figure 2: Forest plot of the effect of multi-component parenting and parental mental health interventions on children's cognitive development

Squares represent the standardised mean difference for each study. The whiskers represent the 95% Cl. The diamond shows the overall pooled effect size using a random-effects model, which is centred at the point estimate, with the diamond width representing the 95% Cl.

intervention content, the Thinking Healthy Program was predominant, focusing on addressing perinatal depressive symptoms by using cognitive behavioural strategies to engage parents. Study characteristics are shown in the table; no study reported the methods used to identify participants' sex or gender.

Studies were done across 14 countries, with six (24%) in the USA. Regarding the mean age of children at baseline, 17 (71%) studies included children with a mean age younger than 12 months, of which ten (59%) studies enrolled pregnant mothers and seven (29%) included children aged 12 months or older (table). Children's mean age-not including those enrolled during pregnancy-was 14.4 months (SD 8.0). 12 of 24 interventions targeted specific populations, such as "at-risk" or "socially vulnerable" families (eg, those with psychosocial vulnerabilities or a combination of socioeconomic and demographic risk factors), adolescent mothers, low-income families, and families with preterm infants. Six of these targeted interventions reported the baseline prevalence of mental health problems (10.3-41.5%). Of the 12 universal interventions, three (25%) reported baseline rates of  $3 \cdot 2-45 \cdot 5\%$  mental health problems. None of the remaining 15 (63%) studies (six [25%] targeted and nine [38%] universal) reported these data.

Mean intervention duration was 14 months (SD 11), ranging from 1 day to 38 months (table). Various techniques were used across interventions to increase participant engagement and support behaviour change. Interventions used a mean of 3.7 behaviour change techniques (SD 2.0), with performance-based techniques such as demonstrations and modelling of key behaviours or direct parent–child engagement being the most widely used (n=13 [15%]; table).

The mean total risk of bias score was  $5 \cdot 00$  (SD  $2 \cdot 73$ ), ranging from 0 to  $11 \cdot 00$  on a 0–14-point scale.<sup>22</sup> Risk of bias was generally low for random sequence generation, blinding of outcome assessment, incomplete outcome data, and selective reporting (appendix p 17). All but one study had a high or unclear risk of bias for the masking of participants and personnel (appendix p 17), a common issue in psychoeducational and behavioural parenting interventions when parents' active involvement in the intervention makes masking difficult. Egger's regression tests were non-significant and funnel plots were fairly symmetrical, suggesting little evidence of publication bias for outcomes included in the metaanalysis (appendix pp 18–19).

Following our preregistered protocol, we did metaanalyses on outcomes reported in at least ten studies. We therefore did not conduct meta-analyses on positive or negative parenting practices and parenting stress in female and male caregivers, and depressive symptoms in male caregivers. In the narrative synthesis, improvements were seen in positive parenting practices but not in negative parenting practices or parenting stress in male and female caregivers (appendix p 20). Only one (out of three) study found significant reductions in depressive symptoms in male caregivers; the others found no significant effect. Data from 23 articles (22 RCTs) were included in the meta-analysis.

11 studies provided 25 effect sizes for children's cognitive development. The Bayley Scales of Infant and Toddler Development was the most frequently used measure (seven [64%] studies). The pooled result showed a small positive effect of multi-component parenting and parental mental health interventions on children's cognitive outcomes (SMD 0.19 [95% CI 0.04-0.34];  $I^2=69\%$ ; figure 2).

20 studies provided 32 effect sizes for children's socialemotional outcomes. The most frequently used measures were the (Brief) Infant and Toddler Social Emotional Assessment and the Ages and Stages Questionnaire (Social Emotional; five [21%] studies each). The pooled result showed a small positive effect on children's socialemotional outcomes (SMD 0.26 [95% CI 0.17-0.34];  $I^2=47\%$ ; figure 3).

16 studies provided 16 effect sizes for depressive symptoms in female caregivers. The Centre for

| Fergusson et al (2005) <sup>38</sup>                                       |      |              | difference (95% CI)  |
|--|------|--------------|----------------------|
| 11   | 1.20 | •            | 0·24 (0·22 to 0·26)  |
| Hayes et al (2008)**   | 0.60 |              | 1.20 (0.28 to 2.12)  |
| Hayes et al (2008)41   | 0.55 | <b>_</b>     | 1.72 (0.76 to 2.68)  |
| Kemp et al (2011) <sup>44</sup>  | 3.73 |              | 0.28 (0.00 to 0.55)  |
| Kemp et al (2011) <sup>44</sup>  | 3.72 | _ <b></b>    | 0·34 (0·07 to 0·62)  |
| Doss et al (2014); <sup>35</sup> Tomfohr-Madsen et al (2020) <sup>36</sup> | 2.64 |              | 0.69 (0.32 to 1.06)  |
| Doss et al (2014); <sup>35</sup> Tomfohr-Madsen et al (2020) <sup>36</sup> | 2.69 | <b>_</b>     | 0.45 (0.09 to 0.82)  |
| Barlow et al (2015) <sup>30</sup>  | 2.92 |              | 0·23 (-0·11 to 0·57) |
| Barlow et al (2015) <sup>30</sup>  | 3.68 |              | 0·23 (-0·05 to 0·51) |
| Castel et al (2016) <sup>33</sup>  | 1.62 | <b>_</b>     | 0.90 (0.38 to 1.41)  |
| Castel et al (2016) <sup>33</sup>  | 1.56 | <b>_</b>     | 1.21 (0.68 to 1.73)  |
| Rockers et al (2018)47   | 5.41 |              | 0.25 (0.08 to 0.42)  |
| Barnhart et al (2020) <sup>31</sup>  | 1.18 |              | 0·35 (-0·27 to 0·98) |
| Maselko et al (2020)45   | 5.06 | _ <b>_</b>   | 0.00 (-0.19 to 0.19) |
| Slade et al (2020)50   | 0.21 | <u>_</u>     | 1.43 (-0.16 to 3.02) |
| Baumgartner et al (2021) <sup>32</sup>                                     | 4.76 | <b>-</b>     | 0·35 (0·15 to 0·56)  |
| Demeusey et al (2021) <sup>34</sup>  | 3.92 |              | 0·13 (-0·13 to 0·39) |
| Demeusey et al (2021) <sup>34</sup>  | 3.92 | _ <b></b>    | 0·11 (-0·15 to 0·37) |
| Luoto et al (2021) <sup>25</sup>   | 5.88 |              | 0·19 (0·05 to 0·34)  |
| Jensen et al (2021) <sup>43</sup>  | 2.07 | <b>_</b>     | 0·10 (-0·34 to 0·54) |
| Pitchik et al (2021) <sup>46</sup>   | 4.92 |              | 0·17 (-0·03 to 0·36) |
| Edward and Hans (2022) <sup>37</sup>                                       | 3.71 |              | 0·16 (-0·12 to 0·43) |
| Edward and Hans (2022) <sup>37</sup>                                       | 3.82 | + <b>-</b> - | 0·19 (-0·08 to 0·45) |
| Edward and Hans (2022) <sup>37</sup>                                       | 3.82 |              | 0·14 (-0·12 to 0·41) |
| Havighurst et al (2022) <sup>40</sup>                                      | 0.24 |              | 0·27 (-1·21 to 1·75) |
| Røhder et al (2022) <sup>48</sup>  | 0.55 |              | 0·26 (-0·70 to 1·22) |
| Fisher et al (2023) <sup>39</sup>  | 6.43 |              | 0·17 (0·05 to 0·28)  |
| Tofail et al (2023) <sup>51</sup>  | 3.87 |              | 0·16 (-0·11 to 0·42) |
| Tofail et al (2023) <sup>51</sup>  | 3.87 |              | 0·13 (-0·13 to 0·39) |
| Tofail et al (2023) <sup>51</sup>  | 3.84 | _ <b></b>    | 0·40 (0·14 to 0·67)  |
| Tofail et al (2023) <sup>51</sup>  | 3.79 |              | 0.59 (0.32 to 0.86)  |
| Tofail et al (2023) <sup>51</sup>  | 3.81 |              | 0.52 (0.25 to 0.78)  |
| Overall (I²=47%; Cochran's Q=68%)  |      | <            | 0·26 (0·17 to 0·34)  |

Figure 3: Forest plot of the effect of multi-component parenting and parental mental health interventions on children's social-emotional outcomes Squares represent the standardised mean difference for each study. The whiskers represent the 95% CI. The diamond shows the overall pooled effect size using a random-effects model, which is centred at the point estimate, with the diamond width representing the 95% CI.

Epidemiologic Study–Depression Scale (eight [38%] studies) and the Edinburgh Postnatal Depression Scale (three [14%] studies) were the most common measures. Only five [29%] studies reported whether measures were validated for use against a gold standard to establish local cutoff scores. The meta-analysis revealed that multi-component parenting and parental mental health interventions did not significantly reduce depressive symptoms in female caregivers (SMD -0.18 [95% CI -0.36 to 0.002]; *I*<sup>2</sup>=86%; figure 4).

We did moderator analyses to explore heterogeneity across pooled estimates. We found no difference in the effect of parental health interventions on children's social-emotional outcomes between high-income countries and LMICs. None of the other characteristics significantly moderated interventions' meta-analytical effects on children's cognitive and social-emotional outcomes or depressive symptoms in female caregivers. These results, as well as the results of sensitivity analyses, are summarised in the appendix (pp 21–22).

## Discussion

This global systematic review and meta-analysis summarises the effect of interventions explicitly targeting parenting behaviours and parental mental health on ECD and parent outcomes. These multi-component interventions had small, positive effects on children's cognitive and social–emotional outcomes but not on depressive symptoms in female caregivers. Narrative results suggested improvements in positive parenting but not negative parenting practices or parenting stress in male and female caregivers. However, there were too few studies examining these outcomes for inclusion in the meta-analysis. We provide, to our knowledge, the first quantitative summary of the effect of multi-component parenting and parental mental health interventions on

|  | Weight<br>(%)           | Standardised mean<br>difference (95% CI) |
|--|-------------------------|--|
|  | 2.71                    | -0.61 (-1.49,0.27)                       |
| Kemp et al (2011) <sup>44</sup>  | 7.52                    | 0.08 (-0.19 to 0.35)                     |
| Doss et al (2014); <sup>35</sup> Tomfohr–Madsen et al (2020) <sup>36</sup> | 6.60                    | -0·21 (-0·57 to 0·15)                    |
| Singla et al (2015) <sup>49</sup>  | 3.84                    | -0.39 (-0.62 to -0.16)                   |
| Castel et al (2016) <sup>33</sup>  | 4.75                    | -0.52 (-1.07 to 0.03)                    |
| Rockers et al (2018)47   | 8-44 —                  | 0.03 (-0.16 to 0.22)                     |
| Barnhart et al (2020) <sup>31</sup>  | 4.97                    | -0.02 (-0.55 to 0.50)                    |
| Maselko et al (2020) <sup>45</sup>   | 8-33                    | -0·11 (-0·30 to 0·09)                    |
| Slade et al (2020)50   | 6.67                    | -0·18 (-0·54 to 0·17)                    |
| Baumgartner et al (2021) <sup>32</sup>                                     | 8.17                    | 0·36 (0·15 to 0·57)                      |
| Demeusey et al (2021) <sup>34</sup>  | 7.65                    | -0·14 (-0·40 to 0·12)                    |
| Luoto et al (2021) <sup>25</sup>   | 8.70                    | 0.09 (-0.05 to 0.24)                     |
| Pitchik et al (2021) <sup>46</sup>   | 6.26                    | -0·10 (-0·30 to 0·10)                    |
| Røhder et al (2022) <sup>48</sup>  | 2.38                    | -0.13 (-1.09 to 0.83)                    |
| Jeong et al (2023) <sup>52</sup>   | 3.40                    | -0·17 (-0·39 to 0·05)                    |
| Tofail et al (2023) <sup>51</sup>  | 7.59                    | -0.43 (-0.69 to -0.16)                   |
| Overall (I²=86%; Cochran's Q=105%)   | $\diamond$              | -0·18 (-0·36 to 0·002)                   |
|  | -1.5 -0.5 0 0.5 1.5 2.5 | i  |

Figure 4: Forest plot of the effect of multi-component parenting and parental mental health interventions on depressive symptoms in female caregivers A negative standardised mean difference represents a favourable outcome as it indicates a reduction in depressive symptoms. Squares represent the standardised mean difference for each study. The whiskers represent the 95% CI. The diamond shows the overall pooled effect size using a random-effects model, which is centred at the point estimate, with the diamond width representing the 95% CI.

child and parent outcomes, highlighting both the promise of such interventions and the gaps that need to be addressed to better realise the potential of such interventions to support young children and their families.

Interventions targeting parenting behaviours or parental mental health alone have mixed evidence of improvements in parenting practices, depressive symptoms, and ECD outcomes across different contexts.<sup>11,14</sup> Although the findings of our review are consistent with findings that parenting interventions in the first 3 years of life, including some that begin during pregnancy, have no effect on depressive symptoms in parents, reviews of parenting interventions including a wider age range (0–17 years) of children have found small-to-moderate effects on different parental mental health outcomes in LMICs and high-income countries.<sup>54,55</sup> Given the increased stress of caring for young children and children's increased vulnerability to their environments, child age is likely to moderate intervention effects on child and parent outcomes.<sup>3</sup>

Our finding that multi-component parenting and parental mental health interventions have no effect on depressive symptoms in parents is consistent with the non-significant effect of parenting interventions on the same outcome,<sup>13,14</sup> suggesting that the expected benefits of integrating content on parenting and parental mental health are not being realised. One potential explanation for this finding is that interventions did not include enough content on mental health to significantly influence parents' symptoms. Although the included studies explicitly targeted parenting behaviours and parental mental health, the few studies that reported the topics of sessions revealed that most focused on content about caring for children's health and development, with only one or two sessions focused on helping parents to manage stress and care for their own wellbeing. It is therefore likely that the mental health content was diluted and thus not sufficient for interventions to significantly affect parental mental health outcomes. Evidence from clinical samples supports the hypothesis that adequate support for parental mental health is needed in combination with parenting interventions to improve mental health and parenting outcomes.56 For example, Ammerman and colleagues<sup>57</sup> found parallel improvements in the depressive symptoms and parenting practices of mothers experiencing clinical depression and coming from a low-income background, who received a home-visiting parenting programme and in-home cognitive behavioural therapy. Another potential explanation for the lack of reduction in depressive symptoms is that most caregivers had low symptom severity at baseline. However, we were unable to test this hypothesis as only nine studies provided these data.

In addition to including little content on parental mental health, the design of included interventions made it difficult to identify independent effects of the parenting and parental mental health components on child and parent outcomes. This lack of clarity in the allocation and dosage of intervention components complicates the assessment of the mental health component's effect and its potential synergistic effects with parenting content. To advance understanding on the effect of improving parental mental health and parenting behaviours—as well as understanding what works and why—intervention studies need to use factorial designs to enable examination of the independent and combined effects of each intervention component.

These multi-component interventions could also have improved ECD outcomes without reducing depressive symptoms in caregivers by other mechanisms.7.8 Jeong and colleagues<sup>58</sup> found that both maternal and paternal stimulation significantly explained the effects of a parenting intervention on children's cognitive and social-emotional development. In addition to parenting practices, a meta-analytical review by Goodman and colleagues<sup>59</sup> found significant bidirectional associations between depression and parenting self-efficacy, suggesting that improving parental self-efficacy could improve child outcomes, even in the presence of depression in mothers.<sup>60</sup> Another potential mechanism is parenting knowledge, as one study in this review found that improvements in maternal knowledge led to improvements in child outcomes.<sup>25</sup> Overall, additional mediation studies are needed to empirically identify the pathways through which multi-component interventions lead to improvements in child outcomes.

Despite testing multiple moderators as sources of heterogeneity in meta-analytical estimates, we did not find any that significantly explained variation in intervention effects on child or parent outcomes. This finding might suggest that the observed effects are consistent across different contexts. However, it is also possible that we did not have sufficient power to detect conditional differences in efficacy given the small number of studies included. For example, although the difference was not significant, the effect of multicomponent interventions on children's social-emotional outcomes in high-income countries was twice as large as that in LMICs. Evidence from a meta-analysis of parenting interventions reported significantly larger effects on child outcomes and parenting practices in LMICs than in high-income countries, suggesting that country income level can moderate the effect of parenting interventions.14 However, other factors such as the nature of child-rearing practices (eg, involvement of other caregivers beyond the immediate family61 and availability of parental leave policies) and quality of mental health services (eg, access to community-based services)62 are likely to play a moderating role beyond a country's income level. For example, given the involvement of other caregivers in caring for young children (primarily, but not exclusively, in LMICs), interventions targeting a child's legal parent who spends little time interacting with the child might not have as large an effect on child outcomes. Although some interventions included caregivers other than mothers and fathers, few assessed outcomes in these caregivers. Key contextual factors should be reported alongside implementation and evaluation data to improve understanding of how different factors might moderate intervention effects.

This study has important strengths, including its quantitative synthesis of the effect of multi-component parenting and parental mental health interventions on child and parent outcomes, an exploration of heterogeneity across key study and implementation characteristics, and a narrative review of implementation features. However, some limitations are worth noting. In addition to the limitations regarding insufficient reported data on the mental health content of included interventions, we were possibly unable to detect some significant effects in the meta-analysis and moderator analysis because of the small number of included studies. For example, although the effect of multi-component parenting and parental mental health interventions on depressive symptoms in female caregivers was not significant, the magnitude of the effect (SMD -0.18) was similar to the effect on child cognitive outcomes (SMD 0.19). Additionally, we were unable to conduct meta-analyses for male caregiver outcomes because of the small number of studies that measured these outcomes. Given increasing evidence of the important contribution of male caregivers to their children's development and family wellbeing, future interventions should include fathers in sessions and measure intervention impacts on fathers' outcomes.63 Most studies originated from high-income countries, thus potentially limiting the applicability of results to LMICs, as there is evidence that the effect of parenting interventions on ECD and parent outcomes are moderated by country income level.14 Studies also used various measures to assess parental mental health outcomes, and information on their psychometric properties was not consistently reported. Results should therefore be interpreted in light of these limitations.

To better understand the effect of multi-component parenting and parental mental health interventions on child and parent outcomes, our findings emphasise a need for an evaluation study design that explicitly examines the independent and joint effects of the same parenting and mental health interventions by testing them in isolation and in combination with one another.<sup>64</sup> This approach can enable a clearer understanding of the structure of these multi-component interventions and their efficacy in improving child and parent outcomes. These insights are crucial to designing interventions that support parents in promoting their own wellbeing and their children's healthy development.

#### Contributors

MNA was responsible for the conceptualisation of the study, with input from AAA, SHG, and JJ, and was involved in screening, extraction, and evaluation of data, in addition to writing the first draft of the paper. MNA ran the meta-analysis. AAA was involved in screening, extraction, and evaluation of data, contributed to writing the first draft of the paper, and was primarily responsible for editing the manuscript on the basis of co-author feedback. PAB prepared and executed the search strategy, with input from MNA and AAA. All authors read, provided input at all stages, and verified the data. MNA had full access to all the data in the study and had final responsibility for the decision to submit for publication.

# Declaration of interests

We declare no competing interests.

#### Data sharing

All data collected for this Article, including the study protocol, data extraction tables, and the statistical analysis, will be available from the publication date. To access these data, please direct requests to the corresponding author.

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#### References

- Lu C, Black MMP, Richter LMP. Risk of poor development in young children in low-income and middle-income countries: an estimation and analysis at the global, regional, and country level. *Lancet Glob Health* 2016; 4: e916–22.
- 2 Black MM, Walker SP, Fernald LCH, et al. Early childhood development coming of age: science through the life course. *Lancet* 2017; 389: 77–90.
- 3 Shonkoff JP, Richter L, van der Gaag J, Bhutta ZA. An integrated scientific framework for child survival and early childhood development. *Pediatrics* 2012; **129**: e460–72.
- 4 Goodman SH, Garber J. Evidence-based interventions for depressed mothers and their young children. *Child Dev* 2017; 88: 368–77.
- 5 Cameron EE, Sedov ID, Tomfohr-Madsen LM. Prevalence of paternal depression in pregnancy and the postpartum: an updated meta-analysis. J Affect Disord 2016; 206: 189–203.
- 6 Gelaye B, Rondon MB, Araya R, Williams MA. Epidemiology of maternal depression, risk factors, and child outcomes in lowincome and middle-income countries. *Lancet Psychiatry* 2016; 3: 973–82.
- 7 Barker B, Iles JE, Ramchandani PG. Fathers, fathering and child psychopathology. *Curr Opin Psychol* 2017; 15: 87–92.
- 8 Goodman SH, Simon HFM, Shamblaw AL, Kim CY. Parenting as a mediator of associations between depression in mothers and children's functioning: a systematic review and meta-analysis. *Clin Child Fam Psychol Rev* 2020; 23: 427–60.
- 9 Sweeney S, MacBeth A. The effects of paternal depression on child and adolescent outcomes: a systematic review. J Affect Disord 2016; 205: 44–59.
- 10 Fang Y, Luo J, Boele M, Windhorst D, van Grieken A, Raat H. Parent, child, and situational factors associated with parenting stress: a systematic review. *Eur Child Adolesc Psychiatry* 2022; 33: 1687–705.
- 11 Goodman SH, Cullum KA, Dimidjian S, River LM, Kim CY. Opening windows of opportunities: evidence for interventions to prevent or treat depression in pregnant women being associated with changes in offspring's developmental trajectories of psychopathology risk. *Dev Psychopathol* 2018; **30**: 1179–96.
- 12 Tsivos ZL, Calam R, Sanders MR, Wittkowski A. Interventions for postnatal depression assessing the mother–infant relationship and child developmental outcomes: a systematic review. *Int J Womens Health* 2015; 7: 429–47.
- 13 Jeong J, Pitchik HO, Yousafzai AK. Stimulation interventions and parenting in low- and middle-income countries: a meta-analysis. *Pediatrics* 2018; 141: 1.
- 14 Jeong J, Franchett EE, Ramos de Oliveira CV, Rehmani K, Yousafzai AK. Parenting interventions to promote early child development in the first three years of life: a global systematic review and meta-analysis. *PLoS Med* 2021; 18: e1003602.
- 15 Fisher J, Cabral de Mello M, Patel V, et al. Prevalence and determinants of common perinatal mental disorders in women in low- and lower-middle-income countries: a systematic review. *Bull World Health Organ* 2012; **90**: 139G–49G.
- 16 Herba CMD, Glover V, Ramchandani PGD, Rondon MBMD. Maternal depression and mental health in early childhood: an examination of underlying mechanisms in low-income and middle-income countries. *Lancet Psychiatry* 2016; 3: 983–92.
- 17 Cui C, Li M, Yang Y, Liu C, Cao P, Wang L. The effects of paternal perinatal depression on socioemotional and behavioral development of children: a meta-analysis of prospective studies. *Psychiatry Res* 2020; 284: 112775.

- 18 Pitchik HO, Chung EO, Fernald LCH. Cross-cultural research on child development and maternal mental health in low- and middleincome countries. *Curr Opin Behav Sci* 2020; 36: 90–97.
- 19 Ahun MN, Ali NB, Hentschel E, Jeong J, Franchett E, Yousafzai AK. A meta-analytic review of the implementation characteristics in parenting interventions to promote early child development. *Ann N Y Acad Sci* 2024; **1533**: 99–144.
- 20 Page MJ, McKenzie JE, Bossuyt PM, et al. The PRISMA 2020 statement: an updated guideline for reporting systematic reviews. *J Clin Epidemiol* 2021; 134: 178–89.
- 21 Rogers A, Obst S, Teague SJ, et al. Association between maternal perinatal depression and anxiety and child and adolescent development: a meta-analysis. JAMA Pediatr 2020; 174: 1082–92.
- 22 Higgins JP, Altman DG, Gøtzsche PC, et al. The Cochrane Collaboration's tool for assessing risk of bias in randomised trials. *BMJ* 2011; 343: d5928.
- 23 Borenstein M. Introduction to meta-analysis, 2nd edn. Hoboken, NJ: John Wiley & Sons, 2021.
- 24 Durlak JA. How to select, calculate, and interpret effect sizes. J Pediatr Psychol 2009; 34: 917–28.
- 25 Luoto JE, Lopez Garcia I, Aboud FE, et al. Group-based parenting interventions to promote child development in rural Kenya: a multiarm, cluster-randomised community effectiveness trial. *Lancet Glob Health* 2021; 9: e309–19.
- 26 Yousafzai AK, Aboud F. Review of implementation processes for integrated nutrition and psychosocial stimulation interventions. *Ann N Y Acad Sci* 2014; 1308: 33–45.
- 27 Aboud FE, Yousafzai AK. Global health and development in early childhood. Annu Rev Psychol 2015; 66: 433–57.
- Pedersen GA, Lakshmin P, Schafer A, et al. Common factors in psychological treatments delivered by non-specialists in low- and middle-income countries: manual review of competencies. *J Behav Cogn Ther* 2020; **30**: 165–86.
- 29 Fisher Z, Tipton E. robumeta: an R-package for robust variance estimation in meta-analysis. arXiv 2015; published online March 7. https://doi.org/10.48550/arXiv.1503.02220 (preprint).
- 30 Barlow A, Mullany B, Neault N, et al. Paraprofessional-delivered home-visiting intervention for American Indian teen mothers and children: 3-year outcomes from a randomized controlled trial. *Am J Psychiatry* 2015; **172**: 154–62.
- Barnhart DA, Farrar J, Murray SM, et al. Lay-worker delivered home visiting promotes early childhood development and reduces violence in Rwanda: a randomized pilot. J Child Fam Stud 2020; 29: 1804–17.
- Baumgartner JN, Ali M, Gallis JA, et al. Effect of a lay counselordelivered integrated maternal mental health and early childhood development group-based intervention in Northern Ghana: a cluster-randomized controlled trial. *Glob Ment Health* 2021; 8: e18.
- 33 Castel S, Creveuil C, Beunard A, Blaizot X, Proia N, Guillois B. Effects of an intervention program on maternal and paternal parenting stress after preterm birth: a randomized trial. *Early Hum Dev* 2016; **103**: 17–25.
- 34 Demeusy EM, Handley ED, Manly JT, Sturm R, Toth SL. Building healthy children: a preventive intervention for high-risk young families. *Dev Psychopathol* 2021; 33: 598–613.
- 35 Doss BD, Cicila LN, Hsueh AC, Morrison KR, Carhart K. A randomized controlled trial of brief coparenting and relationship interventions during the transition to parenthood. J Fam Psychol 2014; 28: 483–94.
- 36 Tomfohr-Madsen LM, Giesbrecht G, Madsen JW, MacKinnon A, Le Y, Doss B. Improved child mental health following brief relationship enhancement and co-parenting interventions during the transition to parenthood. *Int J Environ Res Public Health* 2020; 17: 766.
- 37 Edwards RC, Hans SL. Young mother risk-taking moderates doula home visiting impacts on parenting and toddler social-emotional development. *Dev Psychopathol* 2024; 36: 236–54.
- 38 Fergusson DM, Grant H, Horwood LJ, Ridder EM. Randomized trial of the Early Start program of home visitation. *Pediatrics* 2005; 116: e803–09.
- Fisher J, Tran T, Tran H, et al. Structured, multicomponent, community-based programme for women's health and infant health and development in rural Vietnam: a parallel-group cluster randomised controlled trial. *Lancet Child Adolesc Health* 2023; 7: 311–25.

- 40 Havighurst SS, Kehoe CE, Harley AE, Radovini A, Thomas R. A randomized controlled trial of an emotion socialization parenting program and its impact on parenting, children's behavior and parent and child stress cortisol: Tuning in to Toddlers. *Behav Res Ther* 2022; **149**: 104016.
- 41 Hayes L, Matthews J, Copley A, Welsh D. A randomized controlled trial of a mother–infant or toddler parenting program: demonstrating effectiveness in practice. *J Pediatr Psychol* 2008; 33: 473–86.
- 42 Jacobs F, Easterbrooks MA, Goldberg J, et al. Improving adolescent parenting: results from a randomized controlled trial of a home visiting program for young families. *Am J Public Health* 2016; 106: 342–49.
- 43 Jensen SK, Placencio-Castro M, Murray SM, et al. Effect of a homevisiting parenting program to promote early childhood development and prevent violence: a cluster-randomized trial in Rwanda. *BMJ Glob Health* 2021; 6: e003508.
- 44 Kemp L, Harris E, McMahon C, et al. Child and family outcomes of a long-term nurse home visitation programme: a randomised controlled trial. Arch Dis Child 2011; 96: 533–40.
- 45 Maselko J, Sikander S, Turner EL, et al. Effectiveness of a peerdelivered, psychosocial intervention on maternal depression and child development at 3 years postnatal: a cluster randomised trial in Pakistan. *Lancet Psychiatry* 2020; 7: 775–87.
- 46 Pitchik HO, Tofail F, Rahman M, et al. A holistic approach to promoting early child development: a cluster randomised trial of a group-based, multicomponent intervention in rural Bangladesh. BMJ Glob Health 2021; 6: e004307.
- 47 Rockers PC, Zanolini A, Banda B, et al. Two-year impact of community-based health screening and parenting groups on child development in Zambia: follow-up to a cluster-randomized controlled trial. *PLoS Med* 2018; 15: e1002555.
- 48 Røhder K, Aarestrup AK, Væver MS, Jacobsen RK, Schiøtz ML. Efficacy of a randomized controlled trial of a perinatal adaptation of COS-P in promoting maternal sensitivity and mental wellbeing among women with psychosocial vulnerabilities. *PLoS One* 2022; 17: e027/345.
- 49 Singla DR, Kumbakumba E, Aboud FE. Effects of a parenting intervention to address maternal psychological wellbeing and child development and growth in rural Uganda: a community-based, cluster randomised trial. *Lancet Glob Health* 2015; 3: e458–69.
- 50 Slade A, Holland ML, Ordway MR, et al. Minding the baby: enhancing parental reflective functioning and infant attachment in an attachment-based, interdisciplinary home visiting program. *Dev Psychopathol* 2020; 32: 123–37.
- 51 Tofail F, Islam M, Akter F, et al. An integrated mother–child intervention on child development and maternal mental health. *Pediatrics* 2023; 151 (suppl 2): 151.
- 52 Jeong J, Ahun MN, Gunaratna NS, et al. Effects of engaging fathers and bundling parenting and nutrition interventions on early child development and maternal and paternal parenting in Mara, Tanzania: a factorial cluster-randomized controlled trial. J Child Psychol Psychiatry 2024; 65: 694–709.

- 53 Skeen S, Marlow M, du Toit S, et al. Using WhatsApp support groups to promote responsive caregiving, caregiver mental health and child development in the COVID-19 era: a randomised controlled trial of a fully digital parenting intervention. *Digit Health* 2023; 9: 20552076231203893.
- 54 Beelmann A, Arnold LS, Hercher J. Parent training programs for preventing and treating antisocial behavior in children and adolescents: a comprehensive meta-analysis of international studies. Aggress Violent Behav 2023; 68: 101798.
- 55 Backhaus S, Leijten P, Jochim J, Melendez-Torres GJ, Gardner F. Effects over time of parenting interventions to reduce physical and emotional violence against children: a systematic review and metaanalysis. *EClinicalMedicine* 2023; 60: 102003.
- 56 Goodman SH, Garber J. Evidence-based interventions for depressed mothers and their young children. *Child Dev* 2017; 88: 368–77.
- 57 Ammerman RT, Altaye M, Putnam FW, Teeters AR, Zou Y, Van Ginkel JB. Depression improvement and parenting in lowincome mothers in home visiting. *Arch Womens Ment Health* 2015; 18: 555–63.
- 58 Jeong J, Obradović J, Rasheed M, McCoy DC, Fink G, Yousafzai AK. Maternal and paternal stimulation: mediators of parenting intervention effects on preschoolers' development. J Appl Dev Psychol 2019; 60: 105–18.
- 59 Goodman SH, Simon H, McCarthy L, Ziegler J, Ceballos A. Testing models of associations between depression and parenting selfefficacy in mothers: a meta-analytic review. *Clin Child Fam Psychol Rev* 2022; 25: 471–99.
- 60 Dahlberg M, Nordmyr J, Gunnarsdottir H, Forsman AK. The evidenced effects of early childhood interventions to promote mental health and parenting in the Nordic countries: a systematic review. Int J Ment Health Promot 2023; 25: 505–37.
- 61 Cuartas J, Jeong J, Rey-Guerra C, McCoy DC, Yoshikawa H. Maternal, paternal, and other caregivers' stimulation in low- andmiddle-income countries. *PLoS One* 2020; 15: e0236107.
- 62 Lund C, Brooke-Sumner C, Baingana F, et al. Social determinants of mental disorders and the Sustainable Development Goals: a systematic review of reviews. *Lancet Psychiatry* 2018; 5: 357–69.
- 63 Jeong J, Sullivan EF, McCann JK. Effectiveness of father-inclusive interventions on maternal, paternal, couples, and early child outcomes in low- and middle-income countries: a systematic review. *Soc Sci Med* 2023; 328: 115971.
- 64 Jiang Q, Wang B, Qian Y, et al. Effectiveness of a government-led, multiarm intervention on early childhood development and caregiver mental health: a study protocol for a factorial clusterrandomised trial in rural China. *BMJ Open* 2023; **13**: e076644.